Matching Methods for Causal Inference with Time-Series Cross-Sectional Data

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Quantitative Research Methods Workshop
Yale University

March 7, 2019

Motivation and Overview

- Matching methods have become part of toolkit for social scientists
  1. reduces model dependence in observational studies
  2. provides diagnostics through balance checks
  3. clarifies comparison between treated and control units
- Yet, almost all existing matching methods deal with cross-sectional data
- We propose a matching method for time-series cross-sectional data
  1. create a matched set for each treated observation
  2. refine the matched set via any matching or weighting method
  3. compute the difference-in-differences estimator
- Provide a model-based standard error
- Develop an open-source software package PanelMatch
- Empirical applications:
  - Democracy and economic growth (Acemoglu et al.)
  - Interstate war and inheritance tax (Scheve & Stasavage)

Democracy and Economic Growth

- Acemoglu et al. (2017): an up-to-date empirical study of the long-standing question in political economy

- TSCS data set: 184 countries from 1960 to 2010

- Dynamic linear regression model with fixed effects:

$$Y_{it} = \alpha_i + \gamma_t + \beta X_{it} + \sum_{\ell=1}^{4} \left\{ \rho_\ell Y_{i,t-\ell} + \zeta_\ell^\top \mathbf{Z}_{i,t-\ell} \right\} + \epsilon_{it}$$

  - X_{it}: binary democracy indicator
  - Y_{it}: log real GDP per capita
  - $\mathbf{Z}_{it}$: time-varying covariates (population, trade, social unrest, etc.)

- Sequential exogeneity assumption:

$$\mathbb{E}(\epsilon_{it} \mid \{Y_{it'}\}_{t'=1}^{t-1}, \{X_{it'}\}_{t'=1}^{t}, \{\mathbf{Z}_{it'}\}_{t'=1}^{t-1}, \alpha_i, \gamma_t) = 0$$

- Nickell bias $\rightsquigarrow$ GMM estimation with instruments (Arellano & Bond)

Regression Results

ATE ($\hat{\beta}$)	0.787	0.875	0.666	0.917
	(0.226)	(0.374)	(0.307)	(0.461)
$\hat{\rho}_1$	1.238	1.204	1.100	1.046
	(0.038)	(0.041)	(0.042)	(0.043)
$\hat{\rho}_2$	-0.207	-0.193	-0.133	-0.121
	(0.043)	(0.045)	(0.041)	(0.038)
$\hat{\rho}_3$	-0.026	-0.028	0.005	0.014
	(0.028)	(0.028)	(0.030)	(0.029)
$\hat{\rho}_4$	-0.043	-0.036	0.003	-0.018
	(0.017)	(0.020)	(0.024)	(0.023)
country FE	Yes	Yes	Yes	Yes
time FE	Yes	Yes	Yes	Yes
time trends	No	No	No	No
covariates	No	No	Yes	Yes
estimation	OLS	GMM	OLS	GMM
N	6,336	4,416	6,161	4,245

Treatment Variation Plot

- Regression models does not tell us where the variation comes from

- Estimation of counterfactual outcomes requires comparison between treated and control observations

- Identification strategy:
  - within-unit over-time variation
  - within-time across-units variation

Democracy as the Treatment

Autocracy (Control) Democracy (Treatment)
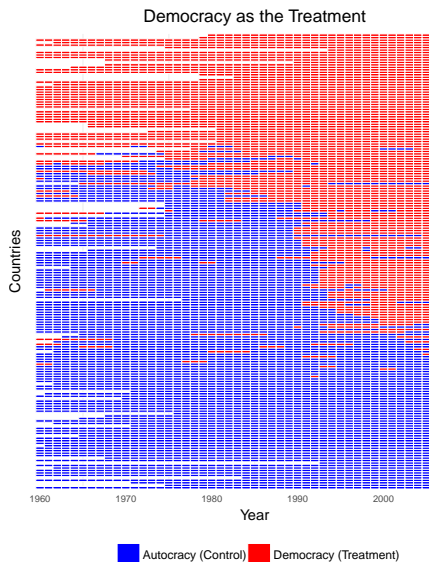
R Package **PanelMatch**

```
──────────────────────────── Data ────────────────────────────
> tail(dem)
     wbcode2 year dem        y   tradewb auth
9379     202 2005   0 589.3235  75.11845    1
9380     202 2006   0 586.1276  78.45123    1
9381     202 2007   0 582.7930  84.64986    1
9382     202 2008   0 563.5891  87.85439    1
9383     202 2009   0 569.2211  97.33238    1
9384     202 2010   0 577.0730 107.33904    1
```

- `wbcode2`: country code
- `year`: year
- `dem`: binary democracy indicator
- `y`: log real GDP per capita (a measure of growth)
- `tradewb`: trade volume
- `auth`: 1-dem

```
──────────────────────────── R Code ────────────────────────────
DisplayTreatment(unit.id = "wbcode2",
                 time.id = "year", legend.position = "none",
                 xlab = "year", ylab = "Country Code", y.size=.5,
                 treatment = "dem", data = dem)
+ theme(axis.text.y = element_blank(), axis.ticks.y = element_blank())
```

War and Taxation

- Inheritance tax plays a central role in wealth accumulations and income inequality
- Scheve and Stasavage (2012): war increases inheritance taxation
- TSCS Data: 19 countries over 185 years from 1816 to 2000
- Static two-way fixed effects model:

$$Y_{it} = \alpha_i + \gamma_t + \beta X_{i,t-1} + \delta^\top \mathbf{Z}_{i,t-1} + \lambda_i t + \epsilon_{it}$$

  - $X_{i,t-1}$: interstate war for country i in year $t-1$
  - Y_{it}: top rate of inheritance tax
  - $\mathbf{Z}_{i,t-1}$: time-varying covariates (leftist executive, a binary variable for the universal male suffrage, and logged real GDP per capita)
  - Strict exogeneity:

$$\mathbb{E}(\epsilon_{it} \mid \mathbf{X}_i, \mathbf{Z}_i, \alpha_i, \gamma_t, \lambda_i) = 0$$

  where $\mathbf{X}_i = (X_{i1}, X_{i2}, \ldots, X_{iT})$ and $\mathbf{Z}_i = (\mathbf{Z}_{i1}^\top, \mathbf{Z}_{i2}^\top, \ldots, \mathbf{Z}_{iT}^\top)^\top$

- Dynamic model without country fixed effects:

$$Y_{it} = \gamma_t + \beta X_{i,t-1} + \rho Y_{i,t-1} + \delta \mathbf{Z}_{i,t-1} + \lambda_i t + \epsilon_{it}$$
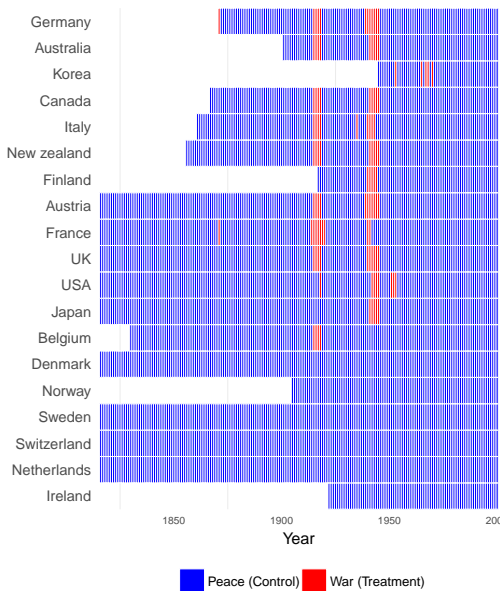
where the strict exogeneity assumption is now given by,

$$\mathbb{E}(\epsilon_{it} \mid \mathbf{X}_i, \mathbf{Z}_i, Y_{i,t-1}, \gamma_t, \lambda_i) = 0$$

- Regression results:

ATE ($\hat{\beta}$)	6.775	1.745	5.970	1.636
	(2.392)	(0.729)	(2.081)	(0.757)
$\hat{\rho}_1$		0.908		0.904
		(0.014)		(0.014)
country FE	Yes	No	Yes	No
time FE	Yes	Yes	Yes	Yes
time trends	Yes	Yes	Yes	Yes
covariates	No	No	Yes	Yes
N	2,780	2,537	2,779	2,536

War as the Treatment

- Treatment is concentrated in a few years

- How should we estimate counterfactual outcomes?

Peace (Control) War (Treatment)

Quantity of Interest and Assumptions

- Choose number of lags $L = 2, \ldots,$ for confounder adjustment
- Choose number of leads, $F = 0, 1, \ldots,$ for short or long term effects
- Average Treatment Effect of Policy Change for the Treated (ATT):

$$\mathbb{E}\left\{ Y_{i,t+F}\left(X_{it} = 1, X_{i,t-1} = 0, \{X_{i,t-\ell}\}_{\ell=2}^{L}\right) - \right.$$
$$\left. Y_{i,t+F}\left(X_{it} = 0, X_{i,t-1} = 0, \{X_{i,t-\ell}\}_{\ell=2}^{L}\right) \mid X_{it} = 1, X_{i,t-1} = 0 \right\}$$

- Assumptions:
  1. No spillover effect
  2. Limited carryover effect (up to L time periods)
  3. Parallel trend after conditioning:

$$\mathbb{E}[Y_{i,t+F}\left(X_{it} = X_{i,t-1} = 0, \{X_{i,t-\ell}\}_{\ell=2}^{L}\right) - Y_{i,t-1}$$
$$\mid X_{it} = 1, X_{i,t-1} = 0, \{X_{i,t-\ell}, Y_{i,t-\ell}\}_{\ell=2}^{L}, \{\mathbf{Z}_{i,t-\ell}\}_{\ell=0}^{L}]$$
$$= \mathbb{E}[Y_{i,t+F}\left(X_{it} = X_{i,t-1} = 0, \{X_{i,t-\ell}\}_{\ell=2}^{L}\right) - Y_{i,t-1}$$
$$\mid X_{it} = 0, X_{i,t-1} = 0, \{X_{i,t-\ell}, Y_{i,t-\ell}\}_{\ell=2}^{L}, \{\mathbf{Z}_{i,t-\ell}\}_{\ell=0}^{L}]$$

Constructing Matched Sets

- Control units with identical treatment history from time $t - L$ to $t - 1$
- Construct a matched set for each treated observation
- Formal definition:

$$\mathcal{M}_{it} = \{i' : i' \neq i, X_{i't} = 0, X_{i't'} = X_{it'} \text{ for all } t' = t - 1, \ldots, t - L\}$$

- Some treated observations have no matched control
  $\rightsquigarrow$ change the quantity of interest by dropping them
- Similar to the risk set of Li et al. (2001) but we do not exclude those who already receive the treatment

An Example of Matched Set

	Country	Year	Democracy	logGDP	Population	Trade
1	Argentina	1974	**1**	888.20	29.11	14.45
2	Argentina	1975	**1**	886.53	29.11	12.61
3	Argentina	1976	**0**	882.91	29.15	12.11
4	Argentina	1977	**0**	888.09	29.32	15.15
5	**Argentina**	**1978**	**0**	**881.99**	**29.57**	**15.54**
6	Argentina	1979	0	890.24	29.85	15.93
7	Argentina	1980	0	892.81	30.12	12.23
8	Argentina	1981	0	885.43	30.33	11.39
9	Argentina	1982	0	878.82	30.62	13.40
10	Thailand	1974	**1**	637.24	43.32	37.76
11	Thailand	1975	**1**	639.51	42.90	41.63
12	Thailand	1976	**0**	645.97	42.44	42.33
13	Thailand	1977	**0**	653.02	41.92	43.21
14	**Thailand**	**1978**	**1**	**660.57**	**41.39**	**42.66**
15	Thailand	1979	1	663.64	40.82	45.27
16	Thailand	1980	1	666.57	40.18	46.69
17	Thailand	1981	1	670.27	39.44	53.40
18	Thailand	1982	1	673.52	38.59	54.22

Find Matched Set using **PanelMatch** Package

Treatment Distribution
Across Units and Time
highlighted matched set for unit id: 177 at time t = 1978 and lag = 4

```
──────────── R Code ────────────
> thailand.1978.index
    <- which(names(msets) == "177.1978")
> thailand.1978.index
[1] 111
> print(msets[111], verbose = TRUE)
$`177.1978`
[1] 6
attr(,"lag")
[1] 4
attr(,"t.var")
[1] "year"
attr(,"id.var")
[1] "wbcode2"
attr(,"treated.var")
[1] "dem"
DisplayTreatment(unit.id = "wbcode2",
                 time.id = "year",
                 legend.position = "none",
                 xlab = "year",
                 ylab = "Country Code",
                 treatment = "dem", data = dem,
                 matched.set = msets[111])
```
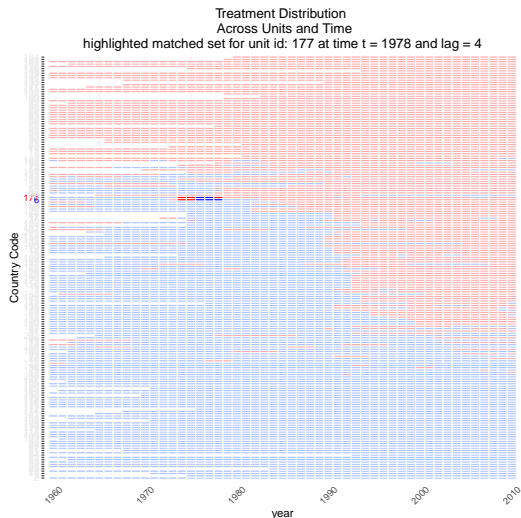
Refining Matched Sets

- Make additional adjustments for past outcomes and confounders
- Use any matching or weighting method
- Mahalanobis distance matching:
  1. Compute the distance between treated and matched control obs.

  $$S_{it}(i') \ = \ \frac{1}{L} \sum_{\ell=1}^{L} \sqrt{(\mathbf{V}_{i,t-\ell} - \mathbf{V}_{i',t-\ell})^{\top} \mathbf{\Sigma}_{i,t-\ell}^{-1} (\mathbf{V}_{i,t-\ell} - \mathbf{V}_{i',t-\ell})}$$

  where $\mathbf{V}_{it'} = (Y_{it'}, \mathbf{Z}_{i,t'+1}^{\top})^{\top}$ and $\mathbf{\Sigma}_{it'} = \mathrm{Cov}(\mathbf{V}_{it'})$
  2. Match the most similar J matched control observations
- Propensity score weighting:
  1. Estimate the propensity score

  $$e_{it}(\{\mathbf{V}_{i,t-\ell}\}_{\ell=1}^{L}) \ = \ \Pr(X_{it} = 1 \mid \{\mathbf{V}_{i,t-\ell}\}_{\ell=1}^{L})$$

  2. Weight each matched control observation by its inverse

```
─────────────── R Code ───────────────
arg.id <- which(names(msets) == "6.1983")

DisplayTreatment(unit.id = "wbcode2",
                 time.id = "year",
                 legend.position = "none",
                 xlab = "year",
                 ylab = "Country Code",
                 treatment = "dem",
                 data = dem,
                 matched.set = msets[arg.id])
```

```
─────────────── R Code ───────────────
pm.obj <- PanelMatch2(lag = 4,
                      time.id = "year",
                      unit.id = "wbcode2",
                      treatment = "dem",
                      outcome = "y",
                      refinement.method =
                        "ps.weight",
                      data = dem,
                      match.missing = T,
                      covs.formula =
                        ~ tradewb +
                        lag("tradewb", 1:4) +
                        lag("y", 1:4))
pm.obj[[arg.1983.idx]]
```


Treatment Distribution
Across Units and Time
highlighted matched set for unit id: 6 at time t = 1983 and lag = 4
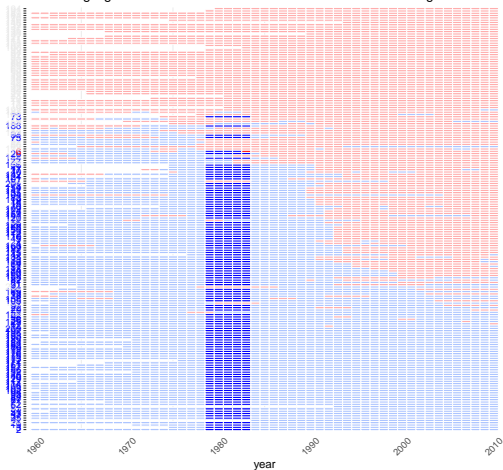
An Example of Refinement

	Country	Year	Democracy	logGDP	Population	Trade	**Weight**
1	Argentina	1979	**0**	890.24	29.85	15.93	1.00
2	Argentina	1980	**0**	892.81	30.12	12.23	1.00
3	Argentina	1981	**0**	885.43	30.33	11.39	1.00
4	Argentina	1982	**0**	878.82	30.62	13.40	1.00
5	**Argentina**	**1983**	**1**	**881.09**	**30.75**	**16.46**	**1.00**
6	Argentina	1984	1	881.76	30.77	15.67	1.00
7	Mali	1979	**0**	542.02	43.80	31.18	0.26
8	Mali	1980	**0**	535.65	43.96	41.82	0.26
9	Mali	1981	**0**	529.10	44.07	41.92	0.26
10	Mali	1982	**0**	522.25	44.45	42.53	0.26
11	**Mali**	**1983**	**0**	**524.84**	**44.74**	**43.65**	**0.26**
12	Mali	1984	0	527.13	44.95	45.92	0.26
13	Chad	1979	**0**	506.71	44.61	44.80	0.27
14	Chad	1980	**0**	498.36	44.84	45.75	0.27
15	Chad	1981	**0**	497.18	45.07	51.58	0.27
16	Chad	1982	**0**	500.07	45.44	43.97	0.27
17	**Chad**	**1983**	**0**	**512.20**	**45.76**	**29.22**	**0.27**
18	Chad	1984	0	511.63	46.04	29.91	0.27
19	Uruguay	1979	**0**	858.39	27.23	41.51	0.47
20	Uruguay	1980	**0**	863.39	27.04	37.99	0.47
21	Uruguay	1981	**0**	864.28	26.93	36.20	0.47
22	Uruguay	1982	**0**	853.36	26.86	35.84	0.47
23	**Uruguay**	**1983**	**0**	**841.87**	**26.83**	**33.36**	**0.47**
24	Uruguay	1984	0	840.08	26.82	42.98	0.47

The Multi-period Difference-in-Differences Estimator

- Compute the weighted average of difference-in-differences among matched control observations
- Weights are based on refinement $\rightsquigarrow$ marginal structural models for the long-term effect of a fixed treatment sequence
- A synthetic control for each treated observation

- The Multi-period DiD estimator:

$$\frac{1}{\sum_{i=1}^{N}\sum_{t=L+1}^{T-F}D_{it}}\sum_{i=1}^{N}\sum_{t=L+1}^{T-F}D_{it}\left\{(Y_{i,t+F}-Y_{i,t-1})-\sum_{i'\in\mathcal{M}_{it}}w_{it}^{i'}(Y_{i',t+F}-Y_{i',t-1})\right\}$$

- Equivalent to the weighted two-way fixed effects estimator:

$$\underset{\beta}{\text{argmin}}\sum_{i=1}^{N}\sum_{t=1}^{T}W_{it}\{(Y_{it}-\overline{Y}_{i}^{*}-\overline{Y}_{t}^{*}+\overline{Y}^{*})-\beta(X_{it}-\overline{X}_{i}^{*}-\overline{X}_{t}^{*}+\overline{X}^{*})\}^{2}$$

where $(\overline{Y}_{i}^{*},\overline{Y}_{t}^{*},\overline{Y}^{*})$ and $(\overline{X}_{i}^{*},\overline{X}_{t}^{*},\overline{X}^{*})$ are weighted averages

Problem of Standard Two-way Fixed Effects Estimator

- In the 2×2 case, the estimator is equivalent to the DiD estimator
- In a more general case, this is not the case
- Instead, the estimator is (approximately) equivalent to:

$$\frac{\frac{1}{N} \sum_{i=1}^{N} S_i^2 (\widehat{Y_i(1)} - \widehat{Y_i(0)}) + \frac{1}{T} \sum_{t=1}^{T} S_t^2 (\widehat{Y_t(1)} - \widehat{Y_t(0)}) - S^2 (\widehat{Y(1)} - \widehat{Y(0)})}{\frac{1}{N} \sum_{i=1}^{N} S_i^2 + \frac{1}{T} \sum_{t=1}^{T} S_t^2 - S^2}$$

where $S_i^2 = \sum_{t=1}^{T} (X_{it} - \overline{X}_i)^2 / (T-1)$, $S_t^2 = \sum_{i=1}^{N} (X_{it} - \overline{X}_t)^2 / (N-1)$, and $S^2 = \sum_{i=1}^{N} \sum_{t=1}^{T} (X_{it} - \overline{X})^2 / (NT-1)$, and for $x = 0, 1$,

$$\widehat{Y_i(x)} = \frac{\sum_{t=1}^{T} \mathbf{1}\{X_{it} = x\} Y_{it}}{\sum_{t=1}^{T} \mathbf{1}\{X_{it} = x\}}, \ \widehat{Y_t(x)} = \frac{\sum_{i=1}^{N} \mathbf{1}\{X_{it} = x\} Y_{it}}{\sum_{i=1}^{N} \mathbf{1}\{X_{it} = x\}},$$

$$\widehat{Y(x)} = \frac{\sum_{i=1}^{N} \sum_{t=1}^{T} \mathbf{1}\{X_{it} = x\} Y_{it}}{\sum_{i=1}^{N} \sum_{t=1}^{T} \mathbf{1}\{X_{it} = x\}}.$$

- We show that appropriate weighting makes this estimator equivalent to the multi-period DiD estimator

Checking Covariate Balance and Computing Standard Error

- Balance for covariate j at time $t - \ell$ in each matched set:

$$B_{it}(j, \ell) = \frac{V_{i,t-\ell,j} - \sum_{i' \in \mathcal{M}_{it}} w_{it}^{i'} V_{i',t-\ell,j}}{\sqrt{\frac{1}{N_1 - 1} \sum_{i'=1}^{N} \sum_{t'=L+1}^{T-F} D_{it'}(V_{i',t'-\ell,j} - \overline{V}_{t'-\ell,j})^2}}$$

- Average this measure across all treated observations:

$$\overline{B}(j, \ell) = \frac{1}{N_1} \sum_{i=1}^{N} \sum_{t=L+1}^{T-F} D_{it} B_{it}(j, \ell)$$

- Standard error calculation $\rightsquigarrow$ consider weight as a covariate
  1. Block bootstrap
  2. Model-based cluster robust standard error within the GMM framework

Empirical Application (1)

- ATT with $L = 4$ and $F = 1, 2, 3, 4$
- We consider democratization and authoritarian reversal
- Examine the number of matched control units
- 18 (13) treated observations have no matched control
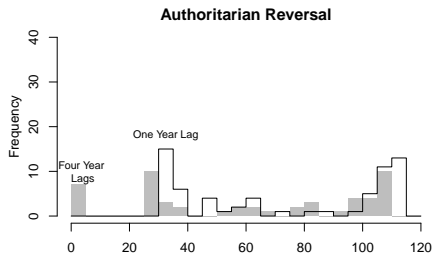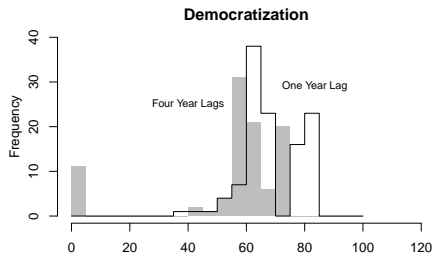
Improved Covariate Balance
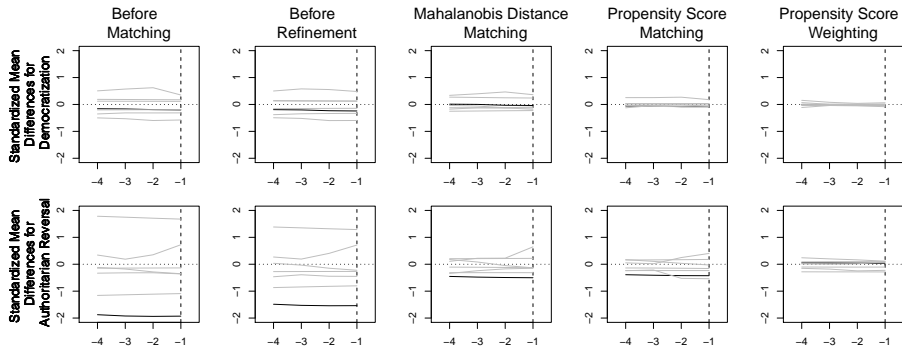
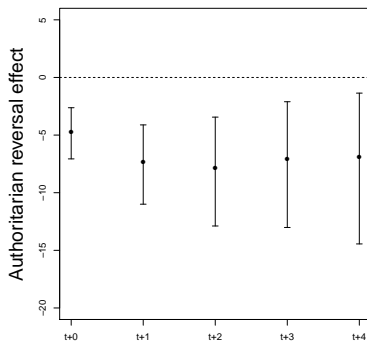# Estimated Causal Effects
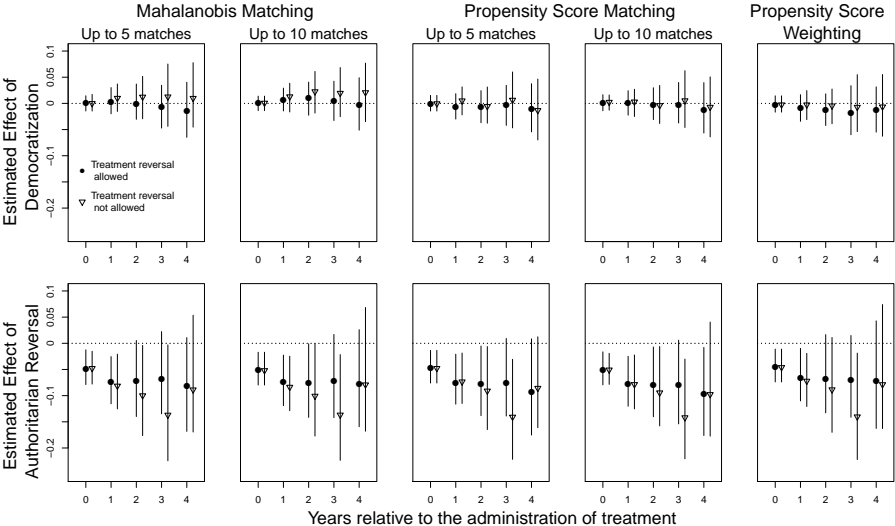
Computing Estimates using **PanelMatch** Package

```R
pm.obj <- PanelMatch2(4, "year", "wbcode2", "auth", "y",
                      refinement.method = "CBPS.weight",
                      data = dem, match.missing = T,
                      covs.formula = ~ tradewb + lag("tradewb", 1:4) +
                      lag("y", 1:4))
res <- PanelEstimate2(lead = 0:4, inference = "bootstrap",
                      qoi = "att", sets = pm.obj, data = dem,
                      outcome.variable = "y")
plot(res,ylim = c(-10, 5), main = "CBPS weighting",
     ylab ="Authoritarian reversal effect")
```
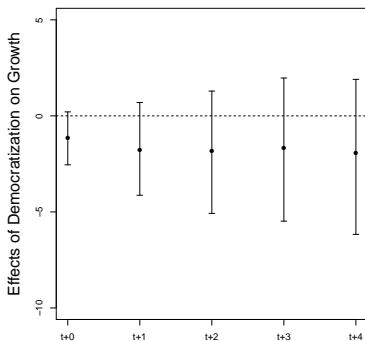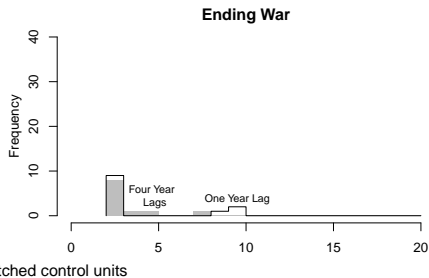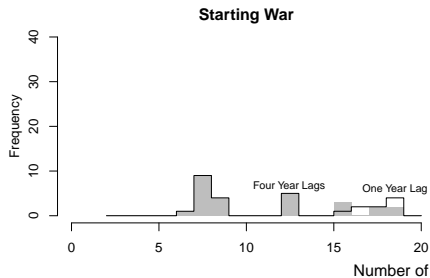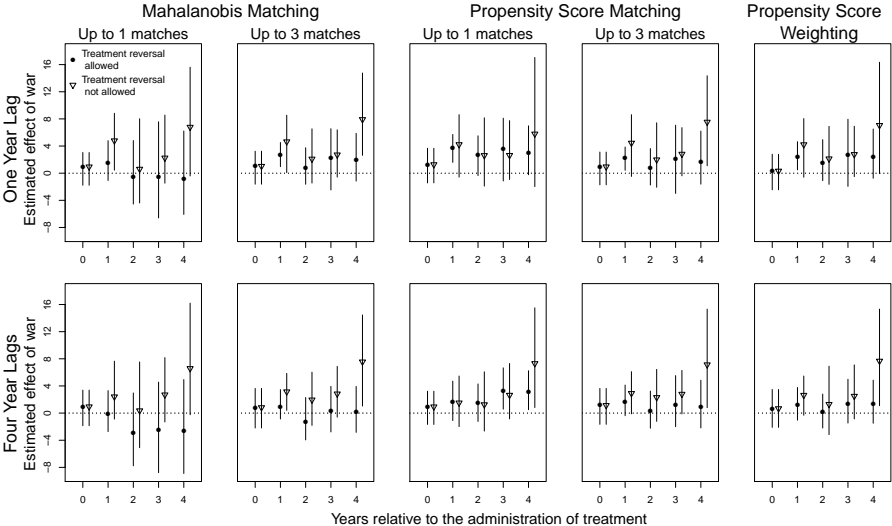
Empirical Application (2)
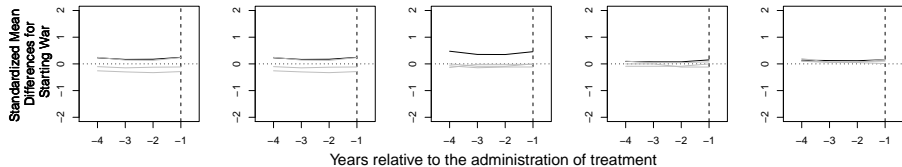
Estimated Causal Effects

Concluding Remarks

- Matching as transparent and simple methods for causal inference
- Yet, matching has not been applied to time-series cross-sectional data

- We propose a matching framework for TSCS data
  1. construct matched sets
  2. refine matched sets
  3. compute difference-in-differences estimator
- Checking covariates and computing standard errors
- R package PanelMatch implements all of these methods

- Ongoing research: addressing possible spillover effects

References

- Available papers:
  1. Imai and Kim. (2019) "When Should We Use Unit Fixed Effects Regression Models for Causal Inference with Longitudinal Data?." *American Journal of Political Science*, Forthcoming."
  2. Imai and Kim. "On the Use of Two-way Fixed Effects Regression Models for Causal Inference with Panel Data."
  3. Imai, Kim, and Wang. "Matching Methods for Causal Inference with Time-Series Cross-Sectional Data."

- Send comments and suggestions to:

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- More information about this and other research:

  https://imai.fas.harvard.edu/