Causal Inference with Spatio-temporal Data: Estimating the Effects of Airstrikes on Insurgent Violence in Iraq

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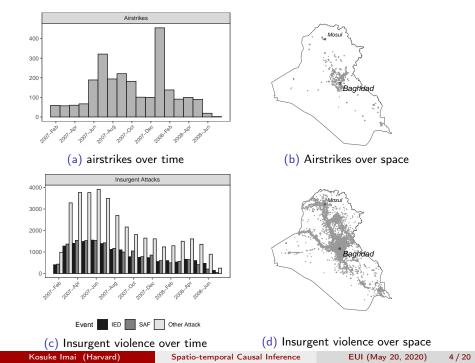
Joint work with Georgia Papadogeorgou (Duke), Jason Lyall (Dartmouth), and Fan Li (Duke)

## Motivation

- Increasing availability of unstructured data in social sciences
  - $\bullet\,$  don't come in a nice matrix form  $\leadsto$  survey, official statistics
  - text, images, audio, video, etc.
- How should we draw causal inference from these new types of data?
- Causal inference with spatio-temporal data
  - a time series of maps as data
  - treatment and outcome event locations in a continuous space
  - applications: crime incidents, disease outbreaks, etc.
- Methodological challenges
  - spillover effects over space
  - 2 carryover effects over time
  - infinitely many possible treatment and outcome locations
- Current practice
  - arbitrary discretization of space
  - assumptions about spillover and carryover effects

## Impacts of Airstrikes on Insurgent Violence in Iraq

- Airstrikes as a principal tool for combating insurgency in civil wars
- Debate: whether or not airstrikes reduce subsequent insurgent attacks (e.g., Kocher *et al.* 2011; Dell and Querubin 2018; Lyall 2019; Mir and Moore 2019)
- Methodological limitations:
  - discretize continuous space into aggregate geographical units
  - simplifying assumptions about spillover and carryover effects
- American air campaign in Iraq:
  - declassified USAF data from Jan. 2007 to July 2008 ("surge" period)
  - daily data with precise geolocation for airstrikes and insurgent attacks
- Drivers of airstrikes:
  - prior patterns of insurgent attacks and airstrikes
  - presence of American forces
  - settlement patterns and road networks
  - economic aid
  - intelligence about high-value targets (small fraction)



## Contributions

- Causal inference with point process treatment and outcome
  - impossible to estimate causal effects of each treatment event
    - unrestricted spillover and carryover effects
    - $\bullet\,$  probability of each treatment realization is zero  $\rightsquigarrow$  lack of overlap
  - stochastic intervention based on the distribution of treatments
  - distribution of airstrikes as a military strategy
- Causal estimands under stochastic intervention
  - expected number of outcome events within a region of interest
  - various stochastic interventions

change the dosage while keeping the distribution identical
change the distribution while keeping the overall dosage constant
intervention over multiple time periods

- The proposed IPW (inverse probability of treatment) estimator
  - overlap and unconfoundedness assumptions
  - consistency and asymptotic normality
- Simulation studies and empirical application

## The Setup

- T time periods:  $t = 1, 2, \ldots, T$
- Treatment variable
  - $\bullet~\Omega:$  set of all possibly infinite locations that can receive the treatment
  - $W_t(s) \in \{0,1\}$ : binary treatment indicator for location s at time t
  - $W_t = \{W_t(s) : s \in \Omega\} \in \mathcal{W}$ : treatment location map at time t
  - $S_{W_t} = \{s \in \Omega : W_t(s) = 1\}$ : set of treatment-active locations at time t with  $|S_{w_t}| < \infty$
  - $\overline{W}_t = (W_1, W_2, \dots, W_t)$ : observed treatment history up to time t
- Outcome variable
  - $Y_t(s)$ ,  $Y_t$ , and  $\overline{Y}_t$  can be similarly defined
  - Potential outcome:  $Y_t(\overline{w}_t)$  where  $w_t \in \mathcal{W}$  is a realized treatment and  $\overline{w}_t = (w_1, w_2, \dots, w_t) \in \mathcal{W}^t$  is a treatment history realization at time t.
  - Observed outcome:  $Y_t = Y_t(\overline{W}_t)$
  - $S_{Y_t(\overline{w}_t)}$ : set of outcome-active locations under treatment history  $\overline{w}_t$
  - History of all potential outcomes up to time t:  $\overline{\mathcal{Y}}_t = \{Y_{t'}(\overline{\boldsymbol{w}}_{t'}) : \overline{\boldsymbol{w}}_{t'} \in \mathcal{W}^{t'}, t' \leq t\}$
- Time-varying confounders:  $X_t$ ,  $\overline{\mathbf{X}}_t$ ,  $X_t(\overline{\mathbf{w}}_{t-1})$ , and  $\overline{\mathcal{X}}_t$

#### Stochastic Intervention

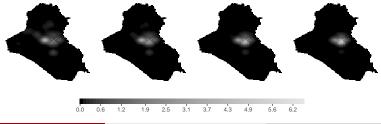
- Stochastic intervention: any distribution of treatment can be used
- We consider Poisson point process  $F_h$ 
  - homogeneous Poisson point process with intensity h: For any disjoint region B<sub>1</sub>, B<sub>2</sub>,..., B<sub>n</sub> ⊂ Ω, the number of events in each region B<sub>i</sub>

$$N_{B_i}(W) \stackrel{\text{indep.}}{\sim} \operatorname{Poisson}(h|B_i|)$$

• non-homogeneous Poisson point process with intensity function  $h(\omega)$ :

$$N_{B_i}(W) \stackrel{\text{indep.}}{\sim} \operatorname{Poisson}\left(\int_{B_i} h(\omega) d\omega\right)$$

• Example:



## Causal Estimands

• Expected number of outcome-active locations in region *B* at time *t* under stochastic intervention *F<sub>h</sub>* conducted at time *t* 

$$\overline{N}_{Bt}(F_h) = \int_{\mathcal{W}} N_B(Y_t(\overline{\boldsymbol{W}}_{t-1}, w_t)) dF_h(w_t)$$

• Further average this quantity over time:

$$\overline{N}_B(F_h) = \frac{1}{T} \sum_{t=1}^T \overline{N}_{Bt}(F_h)$$

• We can compare the different interventions:

$$\tau_B(F_{h'},F_h) = \overline{N}_B(F_{h'}) - \overline{N}_B(F_h)$$

### Stochastic Intervention over Multiple Time Periods

• Consider a non-dynamic stochastic intervention over *M* time periods

$$F_{\boldsymbol{h}} = F_{h_1} \times \cdots \times F_{h_M}$$
 where  $\boldsymbol{h} = (h_1, h_2, \dots, h_M)$ 

• Expected number of outcome-active locations in region B at time t under stochastic intervention  $F_h$  conducted from time t - M + 1 to t

$$\overline{N}_{Bt}(F_{h}) = \int_{\mathcal{W}} \cdots \int_{\mathcal{W}} N_{B}(Y_{t}(\overline{W}_{t-M}, w_{t-M+1}, \dots, w_{t})) dF_{h_{M}}(w_{t-M+1}) \cdots dF_{h_{1}}(w_{t})$$

• Average this quantity over time:

$$\overline{N}_B(F_h) = \frac{1}{T-M+1} \sum_{t=M}^T \overline{N}_{Bt}(F_h)$$

• Comparison of different interventions:

$$\tau_B(F_{h'},F_h) = \overline{N}_B(F_{h'}) - \overline{N}_B(F_h)$$

e.g., lagged effects with  $h_M 
eq h_M'$  and  $\pmb{h}_{-M} = \pmb{h}_{-M}'$ 

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Spatio-temporal Causal Inference

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#### Assumptions

 Unconfoundedness: treatment is independent of all potential (past and future) paths for the outcome and time-varying confounders conditional on the observed history

$$f(W_t \mid \overline{\boldsymbol{W}}_{t-1}, \overline{\boldsymbol{Y}}_{t-1}, \overline{\boldsymbol{X}}_t, \{\overline{\mathcal{Y}}_T, \overline{\mathcal{X}}_T\}) = f(W_t \mid \overline{\boldsymbol{W}}_{t-1}, \overline{\boldsymbol{Y}}_{t-1}, \overline{\boldsymbol{X}}_t)$$

 $\rightsquigarrow$  the generalization of the non-anticipating assumption for time-series experiments (Bojinov and Shephard, 2019)

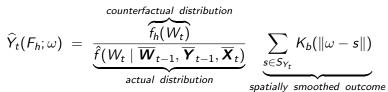
**2** Overlap: there exists a constant  $\delta_W > 0$  such that

$$\underbrace{f(W_t = w \mid \overline{W}_{t-1}, \overline{Y}_{t-1}, \overline{X}_t)}_{\text{propensity score}} > \delta_W \cdot \underbrace{f_h(w)}_{\text{density of } F_h} \quad \text{for all } w \in \mathcal{W}$$

 $\rightsquigarrow$  the ratio  $f_h(w)/f(W_t = w \mid \overline{W}_{t-1}, \overline{Y}_{t-1}, \overline{X}_t)$  is bounded

### The Proposed Estimator

- Inverse probability of treatment weighting (IPW)
- Kernel smoothing of spatial point patterns
- Estimated outcome surface at  $\omega \in \Omega$



where  $K_b$  is the scaled Kernel function with bandwidth parameter b

• Estimated number of outcome-active locations in region B

$$\widehat{\overline{N}}_{Bt}(F_h) = \int_B \widehat{Y}_t(F_h;\omega) d\omega$$

Averaging over time

$$\widehat{\overline{N}}_B(F_h) = \frac{1}{T} \sum_{t=1}^T \widehat{\overline{N}}_{Bt}(F_h)$$

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#### Estimation for Intervention over Multiple Time Periods M

• Estimated outcome surface at  $\omega \in \Omega$ 

$$\widehat{Y}_{t}(F_{\boldsymbol{h}};\omega) = \prod_{j=t-M+1}^{t} \frac{f_{h_{t-j+1}}(W_{j})}{\widehat{f}(W_{j} \mid \overline{\boldsymbol{W}}_{j-1}, \overline{\boldsymbol{Y}}_{j-1}, \overline{\boldsymbol{X}}_{j})} \sum_{s \in S_{Y_{t}}} K_{b}(\|\omega - s\|)$$

where K<sub>b</sub> is the scaled Kernel function with bandwidth parameter b
Estimated number of outcome-active locations in region B

$$\widehat{\overline{N}}_{Bt}(F_{\boldsymbol{h}}) = \int_{B} \widehat{Y}_{t}(F_{\boldsymbol{h}};\omega) d\omega$$

Averaging over time

$$\widehat{\overline{N}}_B(F_h) = \frac{1}{T-M+1} \sum_{t=M}^T \widehat{\overline{N}}_{Bt}(F_h)$$

### Asymptotic Normality and Variance Estimation

• Suppose that there exists v such that

$$\frac{1}{T-M+1}\sum_{t=M}^{T} v_t \stackrel{p}{\longrightarrow} v \quad \text{as } T \longrightarrow \infty$$

where 
$$v_t = \mathbb{V}\left(\widehat{\overline{N}}_{Bt}(F_h) \mid \overline{W}_{t-M}, \overline{\mathcal{Y}}_T, \overline{\mathcal{X}}_T\right).$$

• Then, under some regularity conditions, we have,

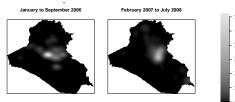
$$\sqrt{T}\left(\widehat{\overline{N}}_{B}(F_{\boldsymbol{h}})-\overline{N}_{B}(F_{\boldsymbol{h}})\right)\overset{d}{\longrightarrow}\mathcal{N}(0,v)$$

where the proof is based on the martingale theory

- Time-specific variance  $v_t$  cannot be estimated
- We use the upper bound  $v_t \leq \mathbb{E}\left(\widehat{\overline{N}}_{Bt}(F_{\boldsymbol{h}})^2 \mid \overline{\boldsymbol{W}}_{t-M}, \overline{\mathcal{Y}}_{\mathcal{T}}, \overline{\mathcal{X}}_{\mathcal{T}}\right)$
- Estimated propensity score ~→ smaller variance

## Empirical Analysis: Setup

• Estimate the baseline treatment distribution  $\phi_0(\omega)$  based on the airstrikes data from January to September, 2006



- How does increasing airstrikes affect insurgent violence?
   → vary c > 0 for h(ω) = c ⋅ φ<sub>0</sub>(ω)
- When long does it take for these effects to be realized?
   → vary c for h<sub>M</sub>(ω) = c · φ<sub>0</sub>(ω) and h<sub>1</sub>(ω) = · · · = h<sub>M-1</sub>(ω) = φ<sub>0</sub>(ω)
- How does the shift in the prioritization of certain locations for airstrikes change the spatial pattern of insurgent attacks?
   → vary α > 0 for h<sub>α</sub>(ω) = c<sub>α</sub> · φ<sub>0</sub>(ω)d<sub>α</sub>(ω) with ∫<sub>Ω</sub> h<sub>α</sub>(ω)dω = c
  - power density  ${\it d}_{lpha}(\omega) \propto {\it d}(\omega)^{lpha}$
  - $d(\omega) =$  the normal density centered at  $s_f$  with precision  $\alpha$

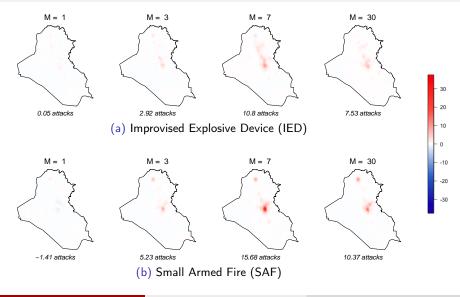
## Propensity Score Model Specification

- Non-homogeneous Poisson point process
  - prior airstrikes over the last day, week, and month

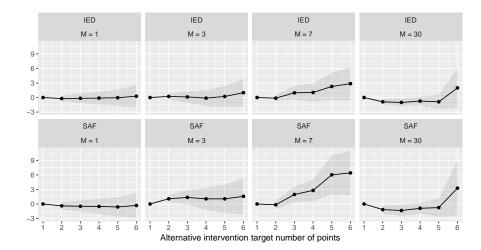
$$\overline{W}_{t-1}^{*}(\omega) = \sum_{j=1}^{7} \sum_{s \in S_{W_{t-j}}} \exp\{-\|s-\omega\|\}$$

- prior insurgent attacks over the last day, week, and month
- prior show-of-force over the last day, week, and month
- amount of US aid in each district over the past month
- distances from major cities, road networks, rivers, and settlements
- log population of governorate (measured in 2003), temporal splines

# Increasing the Expected Number of Airstrikes from 1 to 6 per Day Leads to More Insurgent Attacks

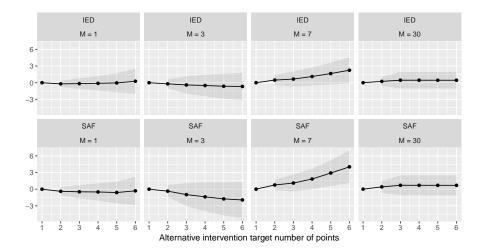


# Effect of Increasing the Airstrikes for M Days on the Number of Insurgent Attacks within Baghdad

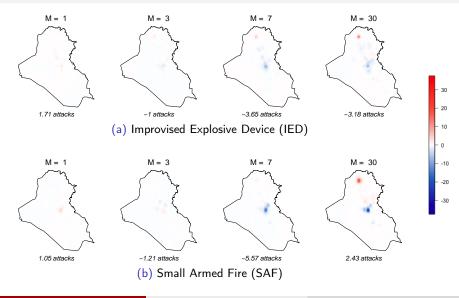


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# Effect of Increasing the Airstrikes *M* **Days Ago** on the Number of Insurgent Attacks within Baghdad



# Increasing the Priority of Baghdad as a Focal Point of Airstrikes Shifts Attacks to Mosul when M is Large



## Concluding Remarks

- A new approach to causal inference with spatio-temporal data
  - directly model point patterns without arbitrary aggregation
  - allow for unstructured spillover and carryover effects
- Key idea: stochastic intervention
  - consider treatment distributions rather than fixed treatment values
  - can handle infinitely many possible treatment locations
  - combine this with spatial smoothing for outcome point process
- Effects of airstrikes on insurgent attacks in Iraq
  - airstrike strategies as stochastic interventions
  - flexible estimation of spillover and carryover effects
- Future research:
  - causal inference with unstructured data such as texts
  - civilian casualty as mediator; comparison with hearts and minds

• Paper at https://imai.fas.harvard.edu/research/spatiotempo.html